

**EC501 Econometric Methods and Applications**

**Problem Set 10: Sketch Solutions**

**Panel Data Models**

1. (a) We begin by stacking the observations over  $t$  for each  $i$ :

$$y_i = X_i\beta + 1_T\alpha_i + \epsilon_i, \quad i = 1, \dots, n,$$

where  $y_i$  is  $T \times 1$ ,  $X_i$  is  $T \times K$ ,  $\alpha_i$  is a scalar and  $1_T$  denotes a  $T \times 1$  vector of ones. We then stack these equations to obtain

$$\begin{pmatrix} y_1 \\ y_2 \\ \vdots \\ y_n \end{pmatrix} = \begin{pmatrix} X_1 \\ X_2 \\ \vdots \\ X_n \end{pmatrix} \beta + \begin{pmatrix} 1_T & 0 & \dots & 0 \\ 0 & 1_T & \dots & 0 \\ & & \vdots & \\ 0 & 0 & \dots & 1_T \end{pmatrix} \begin{pmatrix} \alpha_1 \\ \alpha_2 \\ \vdots \\ \alpha_n \end{pmatrix} + \begin{pmatrix} \epsilon_1 \\ \epsilon_2 \\ \vdots \\ \epsilon_n \end{pmatrix}$$

$$y = X\beta + D\alpha + \epsilon,$$

where  $y$  is  $nT \times 1$ ,  $X$  is  $nT \times K$ ,  $D = I_n \otimes 1_T$  is  $nT \times n$ ,  $\alpha$  is  $n \times 1$  and  $\epsilon$  is  $nT \times 1$ .

- (b) Let  $M_D = I - D(D'D)^{-1}D'$  and note that  $M_DD = 0$ . Then the OLS estimator of  $\beta$  can be obtained by pre-multiplying the model by  $M_D$ :

$$M_D y = M_D X \beta + M_D \epsilon \quad (\text{because } M_D D = 0)$$

$$\Rightarrow y_* = X_* \beta + \epsilon_*$$

$$\Rightarrow b = (X_*' X_*)^{-1} X_*' y_* = (X' M_D X)^{-1} X' M_D y.$$

The estimator of  $\alpha$  can be obtained by noting that the regression residuals,  $e$ , are orthogonal to both  $X$  and  $D$  (i.e.  $X'e = 0$  and  $D'e = 0$ ). Pre-multiplying the residuals  $e = y - Xb - Da$  by  $D'$  results in

$$D'e = D'y - D'Xb - D'Da = 0 \quad \Rightarrow \quad D'Da = D'(y - Xb)$$

$$\Rightarrow a = (D'D)^{-1} D'(y - Xb).$$

- (c) In the expression for  $b$  make the substitution  $y = X\beta + D\alpha + \epsilon$  to obtain

$$b = (X' M_D X)^{-1} X' M_D (X\beta + D\alpha + \epsilon) = \beta + (X' M_D X)^{-1} X' M_D \epsilon,$$

the term involving  $\alpha$  cancelling because  $M_DD = 0$ . It follows that

$$\begin{aligned} E[(b - \beta)(b - \beta)'] &= (X' M_D X)^{-1} X' M_D E(\epsilon\epsilon') M_D X (X' M_D X)^{-1} \\ &= \sigma^2 (X' M_D X)^{-1} \end{aligned}$$

because  $E(\epsilon\epsilon'|x_{it}) = \sigma^2 I_{nT}$  and  $M_D^2 = M_D$ .

(a) The OLS estimates are as follows:

Source	SS	df	MS			
Model	112.705452	3	37.5684839	Number of obs =	90	
Residual	1.33544153	86	.01552839	F( 3, 86) =	2419.34	
				Prob > F =	0.0000	
				R-squared =	0.9883	
				Adj R-squared =	0.9879	
Total	114.040893	89	1.28135835	Root MSE =	.12461	

  

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.8827385	.0132545	66.60	0.000	.8563895	.9090876
x2	.453977	.0203042	22.36	0.000	.4136136	.4943404
lf	-1.62751	.345302	-4.71	0.000	-2.313948	-.9410727
_cons	9.516923	.2292445	41.51	0.000	9.0612	9.972645

(b) The OLS estimates of the fixed effects model are:

Source	SS	df	MS			
Model	113.74827	8	14.2185338	Number of obs =	90	
Residual	.292622872	81	.003612628	F( 8, 81) =	3935.79	
				Prob > F =	0.0000	
				R-squared =	0.9974	
				Adj R-squared =	0.9972	
Total	114.040893	89	1.28135835	Root MSE =	.06011	

  

y	Coef.	Std. Err.	t	P> t	[95% Conf. Interval]	
x1	.9192846	.0298901	30.76	0.000	.8598126	.9787565
x2	.4174918	.0151991	27.47	0.000	.3872503	.4477333
lf	-1.070396	.20169	-5.31	0.000	-1.471696	-.6690963
d1	-.0870617	.0841995	-1.03	0.304	-.2545924	.080469
d2	-.1282976	.0757281	-1.69	0.094	-.2789728	.0223776
d3	-.2959828	.0500231	-5.92	0.000	-.395513	-.1964526
d4	.097494	.0330093	2.95	0.004	.0318159	.1631721
d5	-.063007	.0238919	-2.64	0.010	-.1105443	-.0154697
_cons	9.793004	.2636622	37.14	0.000	9.268399	10.31761

To test the null that all six  $\alpha_i = \alpha$  (i.e. that they are all equal to the same constant value) we have 5 restrictions in all. There are  $K = 3$  regressors, and based on the above estimations we find

$$F = \frac{S_{Pooled} - S_{LSDV}}{S_{LSDV}} \cdot \frac{nT - n - K}{n - 1} = \frac{1.33544153 - 0.292622872}{0.292622872} \cdot \frac{90 - 6 - 3}{5} = 57.732$$

which is distributed as  $F_{5,81}$  under the null. The 5% critical value is approximately 2.34 and hence we reject the null that the  $\alpha_i$  are constant across  $i$  at the 5% level. The same results can be obtained with: `xtreg y x* lf, fe`.

(c) The estimates of the random effects model and the associated LM test are as follows:

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Random-effects GLS regression                Number of obs    =      90
Group variable (i): i                      Number of groups =       6

R-sq:  within = 0.9925                    Obs per group:  min =      15
        between = 0.9856                    avg =             15.0
        overall = 0.9876                    max =             15

Random effects u_i ~Gaussian                Wald chi2(3)     = 11091.33
corr(u_i, X) = 0 (assumed)                  Prob > chi2      =   0.0000

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	y	Coef.	Std. Err.	z	P> z	[95% Conf. Interval]	
	x1	.9066805	.025625	35.38	0.000	.8564565	.9569045
	x2	.4227784	.0140248	30.15	0.000	.3952904	.4502665
	lf	-1.064499	.2000703	-5.32	0.000	-1.456629	-.672368
	_cons	9.627909	.210164	45.81	0.000	9.215995	10.03982
	sigma_u	.12488859					
	sigma_e	.06010514					
	rho	.81193816	(fraction of variance due to u_i)				

Breusch and Pagan Lagrangian multiplier test for random effects:

$$y[i,t] = Xb + u[i] + e[i,t]$$

Estimated results:

	Var	sd = sqrt(Var)
y	1.281358	1.131971
e	.0036126	.0601051
u	.0155972	.1248886

Test: Var(u) = 0

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chi2(1) = 334.85
Prob > chi2 = 0.0000

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The null hypothesis that  $\sigma_u^2 = 0$  is strongly rejected. The additional assumption that is required for this estimator to be consistent is:  $E(u_i|x_{it}) = 0$ .